BMJ Open Effect of smoke-free legislation on respiratory health services use in children with asthma: a populationbased open cohort study in Ontario, Canada

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ABSTRACT

Objective This study will add to existing literature by examining the impact of smoke-free legislation in outdoor areas among children with asthma. We aimed to examine the effect of the 2015 Smoke-Free Ontario Act (SFOA) amendment, which prohibited smoking on patios, playgrounds and sports fields, on health services use (HSU) rates in children with asthma.

Methods We conducted a population-based open cohort study using health administrative data from the province of Ontario, Canada. Each year, all Ontario residents aged 0-18 years with physician diagnosed asthma were included in the study. Annual rates of HSU (emergency department (ED) visits, hospitalisations and physician office visits) for asthma and asthma-related conditions (eq. bronchitis, allergic rhinitis, influenza and pneumonia) were calculated. Interrupted time-series analysis, accounting for seasonality, was used to estimate changes in HSU following the 2015 SFOA.

Results The study population ranged from 618957 individuals in 2010 to 498 812 in 2018. An estimated average increase in ED visits for asthma in infants aged 0-1 years of 0.42 per 100 individuals (95% CI: 0.09 to 0.75) and a 57% relative increase corresponding to the 2015 SFOA was observed. A significant decrease in ED visits for asthma-related conditions of 0.19 per 100 individuals (95% CI: -0.37 to -0.01) and a 22% relative decrease corresponding to the 2015 SFOA was observed. Conclusion Based on the observed positive effect of restricting smoking on patios, playgrounds and sports fields on respiratory morbidity in children with asthma, other jurisdictions globally should consider implementing similar smoke-free policies.

INTRODUCTION

Exposure to environmental tobacco smoke increases the severity of asthma and is an important preventable exposure predisposing asthma exacerbations in children.¹² Secondhand smoke exposure can effectively be reduced by creating comprehensive

Strenghts and limitations of this study

- Using 9 years of health administrative data from Ontario, Canada, our study is one of the largest population-based cohort studies to examine the impact of smoke-free legislation in children with asthma.
- We used interrupted time-series analysis, account-ing for seasonality, to estimate changes in health services use following the 2015 amendment to the Smoke-Free Ontario Act, which prohibited smoking on patios, playgrounds and sports fields.
- This study builds on existing literature in this research area and will help to demonstrate the realworld impact of smoke-free legislation on health outcomes among children and youth with asthma.
- Limitations of this study included lack of clinically relevant risk factors, accounting for reduction in air pollution and smoking due to other campaigns, a control province and inferring individual-level impact of the intervention.

smoke-free public environments enacted through legislation.^{3–5}

While the research on the effects of legislation for smoke-free public environments on adults is well informed, studies on the effects for children are limited and varying. There was no association, immediate nor gradual, between smoke-free legislation and hospital admissions for asthma among children in B Canada and Spain.⁶⁷ In contrast, observational studies in Scotland, England and the USA have found associated reductions in paediatric emergency department (ED) visits for asthma after the implementation of smoke-free legislation.⁸⁻¹⁰ Additionally, in a recent meta-analysis which included some of the studies described previously, smoke-free legislation was found to result in an immediate 9.8% reduction of asthma exacerbations

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requiring hospital attendance and a gradual reduction of 5.9% per year.¹¹ A slightly earlier meta-analysis echoed similar inconsistencies in findings on the effects of legislative smoking bans on respiratory and perinatal health.¹²

In 2015, the Smoke-Free Ontario Act (SFOA) was amended to prohibit smoking on patios, playgrounds and sports fields.¹³ The 2015 SFOA amendment was intended to target and protect children from the harmful effects of smoke exposure by banning smoking in environments where they are often exposed. The present study aims to examine the effect of the 2015 SFOA amendment on rates of asthma and asthma-related health services use (HSU) in children with asthma. This study will contribute evidence on the effects of smoke-free legislation in children and on the effects of this type of smoke-free legislation targeted at reducing smoke exposure among children for the first time. This study will add to existing literature by examining the impact of smoke-free legislation in outdoor areas among children with asthma.

METHODS

Study design and population

We compiled a 9-year population-based open cohort study from January 2010 to December 2018 using health administrative data from the province of Ontario, Canada. Each year, all individuals aged 0-18 years with physician diagnosed asthma and a valid Ontario residence postal code were included in the study. Individuals with a physician diagnosis of asthma were identified using a validated case definition of at least one hospitalisation for asthma or at least two outpatient visits in two consecutive years. This case definition was demonstrated to have 89.7% sensitivity and 70.6% specificity.¹⁴

Data sources

Ontario health administrative databases were used to study HSU over time. The National Ambulatory Care Reporting System captures ED visits, the Canadian Institute for Health Information Discharge Abstract Database captures hospital admissions, the Ontario Health Insurance Plan claims database captures outpatient physician office visits, the provincial Registered Persons Database (RPDB) captures sociodemographic information, and the Ontario Asthma Surveillance Information System captures individuals with asthma. The health administrative databases were linked using unique coded identifiers at ICES (formerly the Institute for Clinical Evaluative Sciences) in Toronto, Ontario.

2015 SFOA amendment

In 2015, the SFOA was amended to prohibit smoking on patios, playgrounds and sports fields to target and protect children from the harmful effects of smoke exposure.¹³ The public was informed of the policy change via provincial and municipal bodies using a variety of strategies ranging from posted signage indicating where smoking is not permitted, news releases,¹⁵ announcements including those streamed live,¹⁶ presentations at public meetings and distribution of information including pamphlets, brochures or web content.¹⁷ Enforcement methods range from implementation of government regulations and inspections for compliance by inspectors or enforcement staff.¹⁷ Individuals who smoke where it is not allowed may be charged with an offence and fined (US\$1000 for a first offence, US\$5000 for any further offence) if convicted.¹³¹⁸

Previously, the SFOA protected Ontarians from exposure to secondhand smoke in enclosed public places and enclosed workplaces since 2006,¹⁹ as well as children under 16 years from such exposure in motor vehicles since 2009.²⁰ Changes around the SFOA since 2015 have focused on protecting youth from the health effects of $\boldsymbol{\xi}$ electronic cigarettes or vaping, such as banning the sale 8 opyright, and supply of electronic cigarettes to youth under 19 years, increasing fines for youth-related sales offences, and restricting the sales and promotion of vapour products in retail stores.^{21 22} including

Statistical analysis

Descriptively, annual rates of HSU (ED visits, hospitalisations and outpatient physician office visits) for asthma and asthma-related conditions (eg, bronchitis, allergic rhinitis, influenza and pneumonia) were calculated. HSU rates were calculated per 100 individuals. Rates were further stratified by age groups as infants (0-1 years), preschoolers (2-4 years), school-aged children (5-9 and ç text 10-14 years) and adolescents (15-18 years). Additional stratifications included sex-based and census-based deprivation quintiles, a proxy measure for socioeconomic status.

atus. We used an interrupted time-series analysis^{23–25} to estimate changes in HSU following the 2015 SFOA amendment. The interrupted time-series analysis controls for baseline level and trend when estimating expected changes in the outcomes of interest that may be attributable to the intervention. The mean monthly HSU rates for asthma and asthma-related conditions in the post-2015 SFOA amendment period were calculated assuming **9** rates without 2015 SFOA and compared with estimated rates with 2015 SFOA.

The characteristics of the study population were summarised based on January 2010, 2015 and 2018. The monthly ED visit, hospitalisation and physician office visit rates were calculated from January 2010 to December lour 2018 for asthma and asthma-related conditions. The 2015 SFOA amendment was effective in January 2015 (month & 61). We used autoregressive models that included autore- \mathbf{g} gressive terms to account for the autocorrelation and seasonality, at a given seasonal lag, in the interrupted time-series data.²⁶ Autocorrelation can also be a consequence of seasonality. The autoregressive terms included adjusted for and balanced any seasonal or cyclical patterns between the preintervention and postintervention periods. The Durbin-Watson statistics were used to test for the presence of autocorrelation of order 12 or smaller. The autoregressive models were estimated using

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the maximum likelihood method. The autocorrelation functions and partial autocorrelation functions were examined in plotted graphs to ascertain that the correlation between any two data points fell within the acceptable range.

The intervention's effects may vary across different subpopulations of individuals. We conducted the interrupted time-series analysis stratifying for these covariates: age group, sex and deprivation quintiles. The absolute effect of the 2015 SFOA amendment was determined as the difference between the estimated HSU rates based on the 2015 SFOA and the calculated counterfactual values at 36 months post-2015 SFOA. The 95% CIs of the absolute difference were calculated according to methods used previously by Zhang et al.²⁷ Additional detail on the calculations is provided in an online supplemental material 1. All statistical analyses were conducted using SAS V.9.4 (SAS Institute).

Patient and public involvement

It was not appropriate or possible to involve patients or the public in the design, or conduct, or reporting, or dissemination plans of our research.

RESULTS

Population characteristics

The study population consisted of Ontario residents 0-18 years of age with physician diagnosed asthma. The study population ranged from 618957 individuals in 2010 to 498812 in 2018 (table 1). Incidence of asthma among individuals 0-18 years of age in Ontario decreased over the study period leading to lower asthma prevalence over time.²⁸ The study population characteristics were similar throughout the 9-year study period. The majority of the study population were male (59.2% in 2010) and the largest proportion were aged 10 to 14 years (31.6%)in 2010). In 2010, the average asthma duration was 8.83 years (SD: 5.19 years) and the average age at asthma diagnosis was 2.94 years (SD: 3.42 years). Most individuals lived in urban areas (90.6% in 2010). The study population was evenly distributed across the five deprivation quintiles, with 18.4% residing in neighbourhoods in the lowest quintile and 22.4% residing in neighbourhoods in the highest quintile in 2010. A small portion of the study population (3.2% in 2010) had immigrated to Canada.

Results from interrupted time-series analysis ED visit rates for asthma

There were no statistically significant changes in the ED visit rates for asthma in the study population post-2015 SFOA except among infants (figure 1A and table 2A). In infants 0-1 years of age, the interrupted time-series analysis showed a significant increase in ED visits for asthma corresponding to the January 2015 SFOA amendment. The estimated average increase in the ED visit rate for asthma in infants was 0.42 per 100 individuals (95% CI: 0.09 to 0.75) from 0.74 to 1.16 per 100 individuals,

pre-2015 and post-2015 SFOA, respectively. This change implied an overall 57% relative increase in asthma ED visits among infants in the 36 months post-2015 SFOA. There were no significant differences observed when stratified by sex and deprivation quintile.

The interrupted time-series analysis showed a significant decrease in ED visits for asthma-related conditions (eg, bronchitis, allergic rhinitis, influenza and pneumonia) corresponding to the 2015 SFOA amendment (figure 1B and table 2B). The estimated average decrease \neg in the ED visit rate for asthma-related conditions was 0.19 per 100 individuals (95% CI: -0.37 to -0.01) from 0.85 to 0.66 per 100 individuals, pre-2015 and post-2015 SFOA, respectively. This change implied a 22% relative decrease in ED visits in the 36 months post-2015 SFOA. The 2 change was more pronounced in infants aged 0-1 years and preschoolers aged 2-4 years (16% and 18% relative decrease in ED visits for asthma-related conditions, respectively). The change was statistically significant in females and those in the highest deprivation quintile (16% and 23% relative decrease in ED visits for asthmarelated conditions, respectively).

Hospital admissions for asthma and asthma-related conditions

From the interrupted time-series analysis, there were no observed statistically significant changes in the hospital admission rates for asthma and asthma-related conditions in the overall study population post-2015 SFOA (figure 2A,B). However, the change in hospital admission for asthma-related conditions was statistically significant in preteens 10–14 years of age and those in the lowest deprivation quintile (43% and 30% relative decrease in hospital admissions for asthma-related conditions, respectively, table 2B). Physician visits for asthma and asthma-related conditions

From the interrupted time-series analysis, there were no statistically significant changes in the physician visit rates for asthma and asthma-related conditions observed in the However, the change in physician visits for asthma-related conditions was statistically in it. 0–1 years of age (20% relative decrease in physician visits for asthma-related conditions, table 2B). There were no significant differences observed when stratified by sex and deprivation quintile.

DISCUSSION

A recent meta-analysis that included studies from North America and Europe found that smoke-free legislation resulted in an immediate and significant reduction of asthma exacerbation requiring hospital attendance.¹¹ However, the extent of this difference, especially in children, remains poorly understood. Some studies found that smoke-free legislation led to reduced paediatric asthma-related HSU,⁸⁻¹⁰ while others found no effect.⁶⁷

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	N or mean or median					
	(% or SD or IQR)					
	2010	2015	2018			
Characteristic	(N=618957)	(N=545950)	(N=498812)			
Sex						
Male	366361 (59.2%)	324160 (59.4%)	296176 (59.4%)			
Female	252 596 (40.8%)	221790 (40.6%)	202636 (40.6%)			
Age group at 1 January of each year (ye	ars)					
0–1	25917 (4.2%)	20742 (3.8%)	15482 (3.1%)			
2–4	64014 (10.3%)	56913 (10.4%)	49301 (9.9%)			
5–9	151 160 (24.4%)	137 438 (25.2%)	126928 (25.4%)			
10–14	195788 (31.6%)	169732 (31.1%)	158997 (31.9%)			
15–18	182078 (29.4%)	161 125 (29.5%)	148104 (29.7%)			
Age at 1 January of each year (years)	, <i>,</i> ,	· · · ·	, <i>,</i> , ,			
Mean (SD)	10.76 (5.06)	10.77 (5.04)	10.90 (4.93)			
Median (IQR)	11.00 (7.00–15.00)	11.00 (7.00–15.00)	11.00 (7.00–15.00)			
Asthma duration (years)	, , , , , , , , , , , , , , , , , , ,	, , , , , , , , , , , , , , , , , , ,	. , ,			
Mean (SD)	8.83 (5.19)	8.88 (5.26)	9.02 (5.14)			
Median (IQR)	8.79 (4.39–13.07)	8.77 (4.35–13.13)	8.85 (4.63–13.18)			
Age at asthma diagnosis (years)						
Mean (SD)	2.94 (3.42)	2.90 (3.30)	2.90 (3.22)			
Median (IQR)	2.00 (0.00–4.00)	2.00 (0.00-4.00)	2.00 (1.00–4.00)			
Rurality	2.00 (0.00					
Urban	560617 (90.6%)	504340 (92.4%)	462566 (92.7%)			
Rural	58340 (9.4%)	41610 (7.6%)	36246 (7.3%)			
Income quintile			00210(11070)			
1 (lowest)	113893 (18.4%)	100531 (18.4%)	88611 (17.8%)			
2	115325 (18.6%)	99354 (18.2%)	89570 (18.0%)			
3	128303 (20.7%)	111529 (20.4%)	102626 (20.6%)			
4	134809 (21.8%)	116679 (21.4%)	110133 (22.1%)			
5 (highest)	124254 (20.1%)	117838 (21.6%)	107 855 (21.6%)			
Missing	2373 (0.4%)	19 (0.0%)	17 (0.0%)			
Deprivation quintile	2010 (0.470)	10 (0.070)	17 (0.070)			
1 (lowest)	114040 (18.4%)	119285 (21.8%)	115566 (23.2%)			
2	120281 (19.4%)	117801 (21.6%)	109217 (21.9%)			
3	119327 (19.3%)	103521 (19.0%)	93180 (18.7%)			
4	121 905 (19.7%)	95335 (17.5%)	84258 (16.9%)			
	138942 (22.4%)	106957 (19.6%)	93933 (18.8%)			
5 (highest) Missing	4462 (0.7%)	3051 (0.6%)	2658 (0.5%)			
	4402 (0.776)	3031 (0.070)	2000 (0.0%)			
Dependency quintile	196 / 70 / 20 10/)	192605 (22 50/)	17/ 585 (25 00/)			
1 (lowest)	186472 (30.1%)	182 695 (33.5%)	174585 (35.0%)			
2	138116 (22.3%)	119893 (22.0%)	108006 (21.7%)			
3	112463 (18.2%)	92291 (16.9%)	81818 (16.4%)			
4 5 (high age)	96476 (15.6%)	80799 (14.8%)	72219 (14.5%)			
5 (highest)	80968 (13.1%)	67221 (12.3%)	59526 (11.9%)			
Missing	4462 (0.7%)	3051 (0.6%)	2658 (0.5%)			

in 0010, 0015 and 0010

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	N or mean or median					
	(% or SD or IQR)	(% or SD or IQR)				
	2010	2015	2018			
Characteristic	(N=618957)	(N=545950)	(N=498812)			
Ethnic concentration quintile						
1 (lowest)	83917 (13.6%)	61236 (11.2%)	52892 (10.6%)			
2	93010 (15.0%)	75770 (13.9%)	66684 (13.4%)			
3	115791 (18.7%)	91276 (16.7%)	82955 (16.6%)			
4	130987 (21.2%)	120077 (22.0%)	111 120 (22.3%)			
5 (highest)	190790 (30.8%)	194540 (35.6%)	182503 (36.6%)			
Missing	4462 (0.7%)	61236 (11.2%) 75770 (13.9%) 91276 (16.7%) 120077 (22.0%) 194540 (35.6%) 3051 (0.6%) 163 605 (30.0%) 108 868 (19.9%) 92 641 (17.0%) 91 223 (16.7%) 86 562 (15.9%) 3051 (0.6%) 18 649 (3.4%) 105 (0.0%) ded, respectively.	2658 (0.5%)			
Residential instability quintile						
1 (lowest)	187 436 (30.3%)	163605 (30.0%)	149482 (30.0%)			
2	128766 (20.8%)	108868 (19.9%)	100204 (20.1%)			
3	110165 (17.8%)	92641 (17.0%)	85906 (17.2%)			
4	105809 (17.1%)	91223 (16.7%)	81809 (16.4%)			
5 (highest)	82319 (13.3%)	86562 (15.9%)	78753 (15.8%)			
Missing	4462 (0.7%)	3051 (0.6%)	2658 (0.5%)			
Immigration						
Yes	19895 (3.2%)	18649 (3.4%)	16301 (3.3%)			
Death count						
Yes	123 (0.0%)	105 (0.0%)	113 (0.0%)			

Using 9years of health administrative data from Ontario, Canada, our study was one of the largest population-based studies to examine the impact of smoke-free legislation in children to date.

This study examined the effect of the 2015 SFOA amendment, which prohibited smoking on patios, playgrounds and sports fields, on HSU rates in children with asthma.¹³ A survey conducted by the Ontario Tobacco Research Unit found significantly decreased self-reported smoking and self-reported secondhand smoking exposure at playgrounds, sports fields, restaurants and park festivals between 2014 and 2015.^{29 30} In the first year after the 2015 SFOA amendment was implemented, compliance with smoking regulations was 96% for patios, 89% for playgrounds and 86% for sporting areas.³⁰ In 2016, compliance with smoking regulations fell to 93% for patios, 73% for playgrounds and 72% for sporting areas.³⁰ Through efforts like the Smoke-Free Ontario Strategy, the smoking rate in Ontario decreased from 20.9% in 2005 to 17.4% in 2014.³¹

Similar to an older study that used health administrative data from Prince Edward Island in Canada, our study also did not find a significant difference post smoke-free

legislation in the asthma-specific hospital admission ≥ rate in children with asthma.⁷ Rather, we found a small increase in ED visits for asthma post-2015 SFOA amend-ment in infants 0–1 years of age. While this difference is statistically significant, the absolute difference is relatively **g** small. Asthma-specific HSU rates are affected by many factors, and it is unclear what may be the contributing S factor for this observed difference. A possible explanation may be confounding by indication where (1) infants 0-1 years of age likely do not spend enough time outdoors in playgrounds and parks to benefit from a smoke-free environment and/or (2) young children being diagnosed by physicians as having asthma tend to have few day-to-day symptoms but are prone to exacerbations,^{32 33} so their risk of having an exacerbation requiring a visit to the ED may not be modified by smoke-free legislation alone.

In this study, we found a statistically significant decrease in ED visits for asthma-related conditions (eg, bronchitis, allergic rhinitis, influenza and pneumonia) corresponding to the 2015 SFOA amendment. Similarly, in a recent meta-analysis, the authors found that the implementation of smoke-free legislation was associated with reductions in rates of hospital attendance for

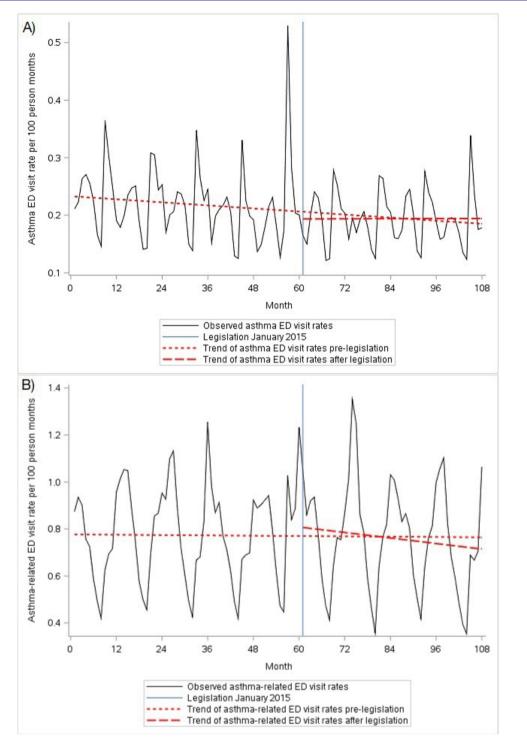


Figure 1 Time series graphs for emergency department (ED) visit rates. (A) Asthma ED visit rate overall. (B) Asthma-related ED visit rate overall.

all respiratory tract infections and for lower respiratory tract infections.¹¹ Our data also showed a significant decrease in asthma-related ED observed in children of lower income families. Compared with children residing in lower deprivation neighbourhoods who may have superior access to private play spaces, children in neighbourhoods with higher deprivation may spend more time in public playgrounds.^{34,35} This may explain the decrease in their asthma-related ED visits before and after the 2015

SFOA amendment. The change may also be attributable to higher baseline exposure to tobacco smoke among children living in lower deprivation neighbourhoods.^{36 37}

Strengths and limitations

This population-based study has several strengths. The use of health administrative data allowed us to collect information about HSU for asthma and asthma-related conditions among all Ontario residents 0–18 years of age

		Mean monthly ra	ate		Relative % chang
Outcome		Without SFOA	With SFOA	Mean monthly change (95% CI)	
A) Effect on he	alth services	use for asthma			
Asthma ED visi	t rate per 100	individuals			
Overall		0.212	0.203	-0.010 (-0.0766 to 0.0576)	-4.48
Age	0–1	0.738	1.158	0.420 (0.0941 to 0.7453)*	56.86
	2–4	0.581	0.582	0.001 (-0.2100 to 0.2121)	0.18
	5–9	0.241	0.216	-0.025 (-0.0677 to 0.0181)	-10.30
	10–14	0.119	0.108	-0.011 (-0.0411 to 0.0192)	-9.22
	15–18	0.100	0.095	-0.006 (-0.0359 to 0.0243)	-5.75
Sex	Male	0.224	0.212	-0.012 (-0.0865 to 0.0633)	-5.19
	Female	0.198	0.188	-0.010 (-0.0595 to 0.0402)	-4.88
Deprivation	Q1 (least)	0.172	0.171	-0.001 (-0.0497 to 0.0487)	-0.29
1	Q2	0.183	0.178	-0.006 (-0.0574 to 0.0464)	-3.00
	Q3	0.207	0.189	0.000 (-0.0781 to 0.0407)	0.00
	Q4	0.225	0.215	-0.010 (-0.0941 to 0.0740)	-4.48
	Q5 (most)	0.283	0.247	-0.036 (-0.1124 to 0.0397)	-12.83
Asthma hospit	()	ate per 100 individua		-0.000 (-0.1124 to 0.0007)	-12.00
Overall		0.056	0.053	–0.003 (–0.0305 to 0.0251)	-4.88
	0–1	0.517	0.761	0.243 (-0.0385 to 0.5254)	47.09
Age				· · · · ·	
	2-4	0.193	0.189	-0.004 (-0.1060 to 0.0984)	-1.97
	5-9	0.057	0.032	-0.026 (-0.0727 to 0.0211)	-45.00
	10–14	0.015	0.012	-0.002 (-0.0115 to 0.0068)	-16.18
-	15–18	0.008	0.008	0.000 (-0.0027 to 0.0033)	3.93
Sex	Male	0.057	0.056	-0.001 (-0.0316 to 0.0292)	-2.13
	Female	0.054	0.049	-0.005 (-0.0297 to 0.0205)	-8.49
Deprivation	Q1 (least)	0.044	0.037	-0.007 (-0.0272 to 0.0136)	–15.41
	Q2	0.047	0.044	-0.003 (-0.0265 to 0.0205)	-6.45
	Q3	0.052	0.051	-0.001 (-0.0287 to 0.0260)	-2.66
	Q4	0.059	0.062	0.002 (-0.0305 to 0.0351)	3.92
	Q5 (most)	0.076	0.071	-0.005 (-0.0451 to 0.0347)	-6.82
Asthma physic	an visit rate p	er individual			
Overall		0.036	0.038	0.002 (–0.0022 to 0.0056)	4.75
Age	0–1	0.212	0.224	0.012 (-0.0245 to 0.0491)	5.80
	2–4	0.099	0.102	0.003 (-0.0062 to 0.0119)	2.90
	5–9	0.045	0.042	-0.004 (-0.0121 to 0.0045)	-8.39
	10–14	0.021	0.022	0.001 (-0.0015 to 0.0038)	5.52
	15–18	0.011	0.012	0.001 (-0.0002 to 0.0023)	10.19
Sex	Male	0.037	0.039	0.002 (-0.0023 to 0.0057)	4.63
	Female	0.036	0.037	0.002 (-0.0021 to 0.0056)	4.90
Deprivation	Q1 (least)	0.037	0.034	-0.004 (-0.0130 to 0.0055)	-10.12
	Q2	0.037	0.034	-0.003 (-0.0104 to 0.0049)	-7.50
	Q3	0.037	0.035	-0.002 (-0.0069 to 0.0038)	-4.31
	Q4	0.037	0.040	0.003 (-0.0022 to 0.0081)	7.88
	Q5 (most)	0.039	0.042	0.003 (-0.0008 to 0.0058)	6.39

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		Mean monthly ra	ate		Relative % change
Outcome		Without SFOA	With SFOA	Mean monthly change (95% CI)	
Asthma-related	d ED visit rate	per 100 individuals			
Overall		0.854	0.664	-0.190 (-0.3722 to -0.0075)*	-22.23
Age	0–1	6.872	5.781	-1.091 (-2.0908 to -0.0908)*	-15.87
	2–4	2.213	1.821	-0.392 (-0.7775 to -0.0063)*	-17.71
	5–9	0.677	0.595	-0.082 (-0.2008 to 0.0366)	-12.13
	10–14	0.352	0.317	-0.035 (-0.1270 to 0.0571)	-9.92
	15–18	0.435	0.369	-0.066 (-0.2024 to 0.0710)	-15.12
Sex	Male	0.792	0.704	-0.088 (-0.1980 to 0.0222)	-11.10
	Female	0.843	0.709	–0.133 (–0.2506 to –0.0159)*	-15.81
Deprivation	Q1 (least)	0.632	0.588	-0.044 (-0.1321 to 0.0439)	-6.97
	Q2	0.762	0.599	-0.163 (-0.3466 to 0.0198)	-21.43
	Q3	0.809	0.702	0.000 (-0.2164 to 0.0031)	0.00
	Q4	0.898	0.747	-0.152 (-0.3439 to 0.0405)	-16.89
	Q5 (most)	1.041	0.804	-0.237 (-0.4514 to -0.0226)*	-22.78
Asthma-related	d hospital adm	nission rate per 100 i	ndividuals		
Overall		0.090	0.073	-0.016 (-0.0388 to 0.0059)	-18.40
Age	0–1	1.642	1.479	-0.163 (-0.9228 to 0.5254)	-9.92
	2–4	0.220	0.174	-0.046 (-0.0966 to 0.0037)	-21.10
	5–9	0.049	0.041	-0.008 (-0.0290 to 0.0128)	-16.45
	10–14	0.022	0.013	–0.009 (–0.0183 to –0.0006)*	-42.62
	15–18	0.016	0.014	-0.002 (-0.0066 to 0.0023)	-13.63
Sex	Male	0.090	0.075	-0.015 (-0.0380 to 0.0079)	-16.76
	Female	0.090	0.071	-0.019 (-0.0425 to 0.0043)	-21.31
Deprivation	Q1 (least)	0.076	0.053	-0.023 (-0.0457 to -0.0006)*	-30.39
	Q2	0.076	0.062	-0.013 (-0.0306 to 0.0044)	-17.32
	Q3	0.089	0.068	-0.021 (-0.0440 to 0.0022)	-23.57
	Q4	0.091	0.083	-0.009 (-0.0275 to 0.0103)	-9.42
	Q5 (most)	0.112	0.100	-0.012 (-0.0437 to 0.0194)	-10.87
Asthma-related	d physician vis	sit rate per individual			
Overall		0.093	0.090	-0.002 (-0.0120 to 0.0072)	-2.59
Age	0–1	0.395	0.315	-0.081 (-0.1561 to -0.0053) *	-20.40
	2–4	0.185	0.177	-0.008 (-0.0260 to 0.0095)	-4.48
	5–9	0.105	0.098	0.107 (-0.0200 to 0.0063)	101.55
	10–14	0.066	0.067	-0.007 (-0.0098 to 0.0104)	-10.34
	15–18	0.056	0.058	0.001 (-0.0067 to 0.0096)	2.58
Sex	Male	0.089	0.087	-0.002 (-0.0116 to 0.0068)	-2.69
	Female	0.098	0.096	-0.002 (-0.0129 to 0.0079)	-2.54
Deprivation	Q1 (least)	0.087	0.077	-0.009 (-0.0245 to 0.0060)	-10.65
	Q2	0.091	0.084	-0.007 (-0.0183 to 0.0045)	-7.64
	Q3	0.094	0.089	-0.006 (-0.0157 to 0.0044)	-6.01
	Q4	0.099	0.099	0.000 (-0.0104 to 0.0110)	0.26
	Q5 (most)	0.099	0.100	0.001 (-0.0097 to 0.0113)	0.81

Continued

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Table 2 Continue	d
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	 Mean monthly rate			Relative % change
Outcome	Without SFOA	With SFOA	Mean monthly change (95% CI)	

ED visit rates for asthma-related conditions.

*Indicates statistical significance.

ED, emergency department; SFOA, smoke-free Ontario act.

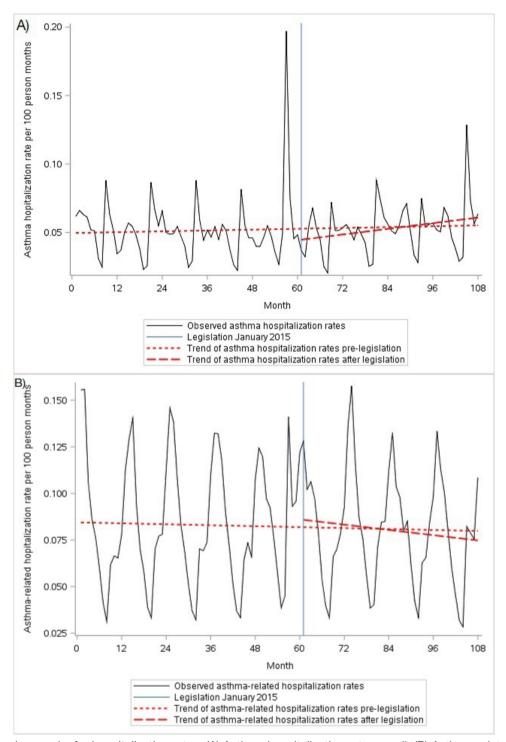


Figure 2 Time series graphs for hospitalisation rates. (A) Asthma hospitalisation rate overall. (B) Asthma-related hospitalisation rate overall.

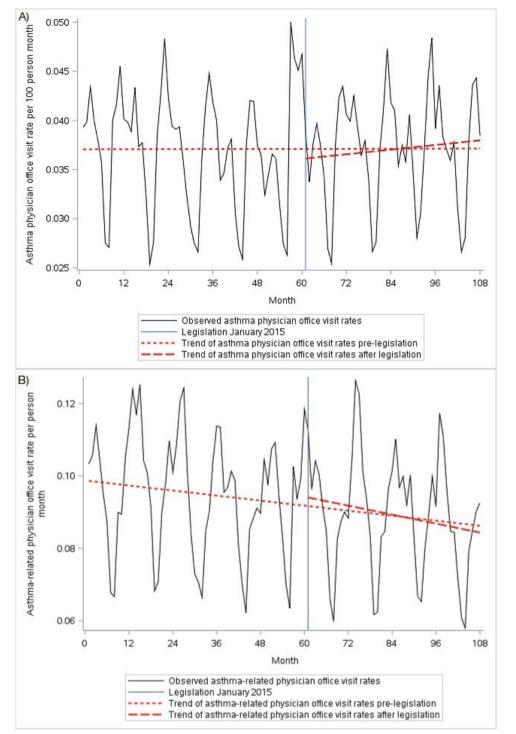


Figure 3 Time series graphs for physician office visit rates. (A) Asthma physician office visit rate overall. (B) Asthma-related physician office visit rate overall.

with asthma over a 9-year period. This data source allowed us to study HSU without selection or response bias, and may allow us to generalise our results to other populations in other regions. There are also some limitations to our study. The use of health administrative data limited our ability to examine clinically relevant risk factors such as asthma severity, medication use, comorbid conditions, parental smoking and unplanned versus planned healthcare visits. Our ability to account for reduction in air pollution and smoking in the population due to other campaigns was also limited. We did not include a control province as smoke-free legislation varies across Canada and identifying a province where such legislation has not been implemented is not straightforward. Furthermore, since the interrupted time-series analysis was used to calculate mean monthly HSU rates on an ecological level, we were limited in our ability to infer the impact of the intervention on an individual level.

CONCLUSION

This population-based study used 9 years of health administrative data to study the impact of the 2015 SFOA amendment on respiratory HSU in children with asthma. We found that while the 2015 SFOA amendment did not impact healthcare use for asthma specifically, it had a greater and significant impact on healthcare use for asthma-related conditions (eg, bronchitis, allergic rhinitis, influenza and pneumonia) in children living with asthma. Based on the positive effect of restricting smoking on patios, playgrounds, and sports fields on respiratory morbidity in children with asthma observed in this study, other jurisdictions globally should consider implementing similar smoke-free policies. Additionally, Ontario should improve compliance with this smoke-free policy by continuing to educate the public on the harmful effects of smoking, especially on children.

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Contributors TT initiated and designed the study, interpreted findings and drafted the manuscript. IF drafted the manuscript, summarised findings and composed tables and figures. JZ conducted all statistical analysis and drafted the manuscript. RM initiated the study and drafted and reviewed the manuscript. KZ and ET conducted a search of the literature, summarised relevant study findings and reviewed the manuscript as submitted and agreed to be accountable for all aspects of the work.

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Competing interests None declared.

Patient consent for publication Not required.

Ethics approval Ethics approval exemption was obtained from the Hospital for Sick Children Research Ethics Board. Ethics approval was exempted because this study uses only health administrative data housed at ICES under the Personal Health Information Protection Act (PHIPA) section 45, for the purpose of analysis or compiling statistical information with respect to the management of, evaluation or monitoring of, the allocation of resources to or planning for all or part of the health system, including the delivery of services, as determined by ICES Privacy Impact Assessment review. Individual consent was not required as only deidentified provincial health administrative data were used in this study.

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Data availability statement Data may be obtained from a third party and are not publicly available. No additional data available.

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